

Mixed Models for Longitudinal Binary Outcomes

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Hedeker, D. (2005). Generalized linear mixed models. In B. Everitt & D. Howell (Eds.), *Encyclopedia of Statistics in Behavioral Science*. Wiley.

Hedeker, D. & Gibbons, R.D. (2006). *Longitudinal Data Analysis*, chapter 9. Wiley.

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Mixed-effects models for categorical outcomes

- dichotomous outcomes
 - mixed-effects logistic regression
- ordinal outcomes
 - mixed-effects ordinal logistic regression
 - * proportional odds model
 - * partial or non-proportional odds model
- nominal outcomes
 - mixed-effects nominal logistic regression
- discrete or grouped time-to-event data
 - mixed-effects dichotomous or ordinal regression
 - * complementary log-log link for proportional (and non-proportional) hazards models

Logistic Regression Model

$$\log \left[\frac{P(Y_i = 1)}{1 - P(Y_i = 1)} \right] = \mathbf{x}'_i \boldsymbol{\beta}$$

- Dichotomous outcome ($Y = 0$ absence, $Y = 1$ presence).
- Function that links probabilities to regressors is the logit (or log odds) function $\log [P/(1 - P)]$. Logit is called the link function.

The model can be written in terms of probabilities:

$$P(Y_i = 1) = \frac{1}{1 + \exp(-\mathbf{x}'_i \boldsymbol{\beta})}$$

- Model is a linear model for the logits, not for the probabilities. Logits can take on any values between negative and positive infinity, probabilities can only take on values between 0 and 1

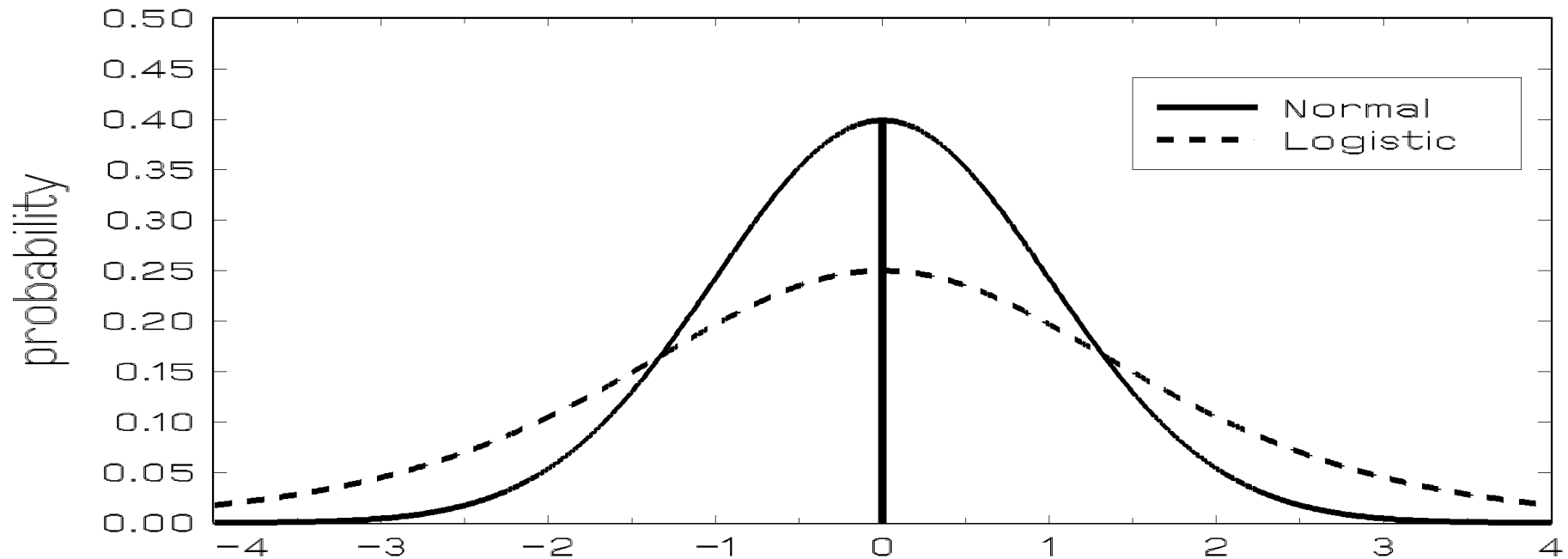
Dichotomous Response and Threshold Concept

Continuous y_i - an unobservable latent variable - related to dichotomous response Y_i via “threshold concept”

Response occurs ($Y_i = 1$) if $\gamma < y_i$

otherwise, a response does not occur ($Y_i = 0$)

Latent Distribution: Normal and Logistic pdf



The Threshold Concept in Practice

“How was your day?” (what is your satisfaction level today?)

- Satisfaction may be continuous, but we usually emit a dichotomous response:



Great Day!



a day ...

Model for Latent Continuous Responses

Consider the model with p covariates for the latent response strength y_i ($i = 1, 2, \dots, N$):

$$y_i = \mathbf{x}'_i \boldsymbol{\beta} + \varepsilon_i$$

- probit: $\varepsilon_i \sim$ standard normal (mean=0, variance=1)
- logistic: $\varepsilon_i \sim$ standard logistic (mean=0, variance= $\pi^2/3$)

$\Rightarrow \boldsymbol{\beta}$ estimates from logistic regression are larger (in abs. value) than from probit regression by approximately $\sqrt{\pi^2/3} = 1.8$

Underlying latent variable

- useful way of thinking of the problem
- not an essential assumption of the model

Random-intercept Logistic Regression Model

Consider the model with p covariates for the response Y_{ij} for subject i ($i = 1, 2, \dots, N$) at time j ($j = 1, 2, \dots, n_i$):

$$\log \left[\frac{P(Y_{ij} = 1 \mid v_{0i})}{1 - P(Y_{ij} = 1 \mid v_{0i})} \right] = \mathbf{x}'_{ij} \boldsymbol{\beta} + v_{0i}$$

where

Y_{ij} = dichotomous response for subject i at time j

\mathbf{x}_{ij} = $(p + 1) \times 1$ covariate vector (includes 1 for intercept)

$\boldsymbol{\beta}$ = $(p + 1) \times 1$ vector of unknown parameters

v_{0i} = subject effects distributed $\mathcal{NID}(0, \sigma_v^2)$

Model for Latent Continuous Responses

Consider the model with p covariates for the $n_i \times 1$ latent response strength y_{ij} :

$$y_{ij} = \mathbf{x}'_{ij}\boldsymbol{\beta} + \nu_{0i} + \varepsilon_{ij}$$

where assuming

- $\varepsilon_{ij} \sim$ standard normal (mean 0 and $\sigma^2 = 1$) leads to mixed-effects probit regression
- $\varepsilon_{ij} \sim$ standard logistic (mean 0 and $\sigma^2 = \pi^2/3$) leads to mixed-effects logistic regression

Underlying latent variable

- not an essential assumption of the model
- useful for obtaining intra-class correlation (r)

$$r = \frac{\sigma_v^2}{\sigma_v^2 + \sigma^2}$$

and for design effect (d)

$$d = \frac{\sigma_v^2 + \sigma^2}{\sigma^2} = 1/(1 - r)$$

ratio of actual variance to the variance that would be obtained by simple random sampling (holding sample size constant)

Scaling of regression coefficients

β estimates from mixed-effects model are larger (in abs. value) than from fixed-effects model by approximately

$$\sqrt{d} = \sqrt{\frac{\sigma_v^2 + \sigma^2}{\sigma^2}}$$

because

- $V(y) = \sigma_v^2 + \sigma^2$ in mixed-effects model
- $V(y) = \sigma^2$ in fixed-effects model

difference depends on size of random-effects variance σ_v^2

Treatment-Related Change Across Time

Data from the NIMH Schizophrenia collaborative study on treatment related changes in overall severity. IMPS item 79, *Severity of Illness*, was scored as:

- 1 = normal
- 2 = borderline mentally ill
- 3 = mildly ill

- 4 = moderately ill
- 5 = markedly ill
- 6 = severely ill
- 7 = among the most extremely ill

The experimental design and corresponding sample sizes:

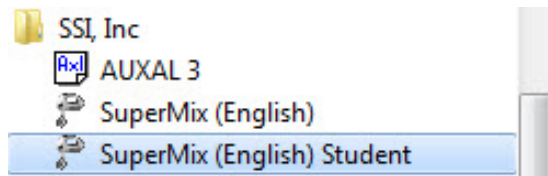
Group	Sample size at Week							<i>completers</i>
	0	1	2	3	4	5	6	
PLC (n=108)	107	105	5	87	2	2	70	65%
DRUG (n=329)	327	321	9	287	9	7	265	81%

Drug = Chlorpromazine, Fluphenazine, or Thioridazine

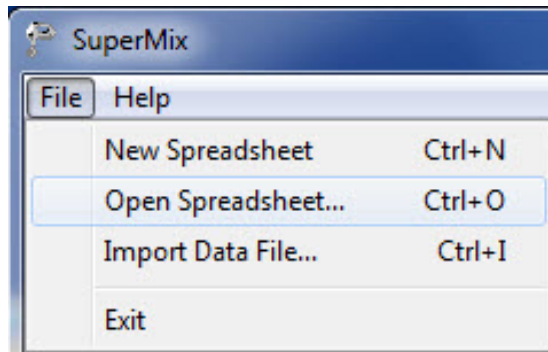
Main question of interest:

- Was there differential improvement for the drug groups relative to the control group?

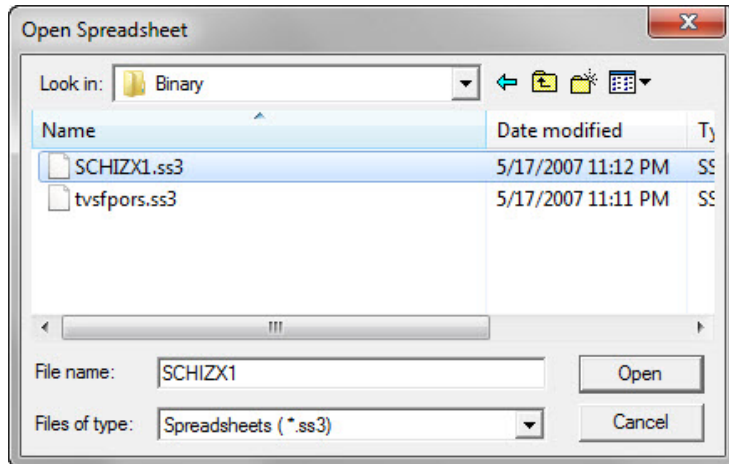
- Under SSI, Inc > “SuperMix (English)” or “SuperMix (English) Student”



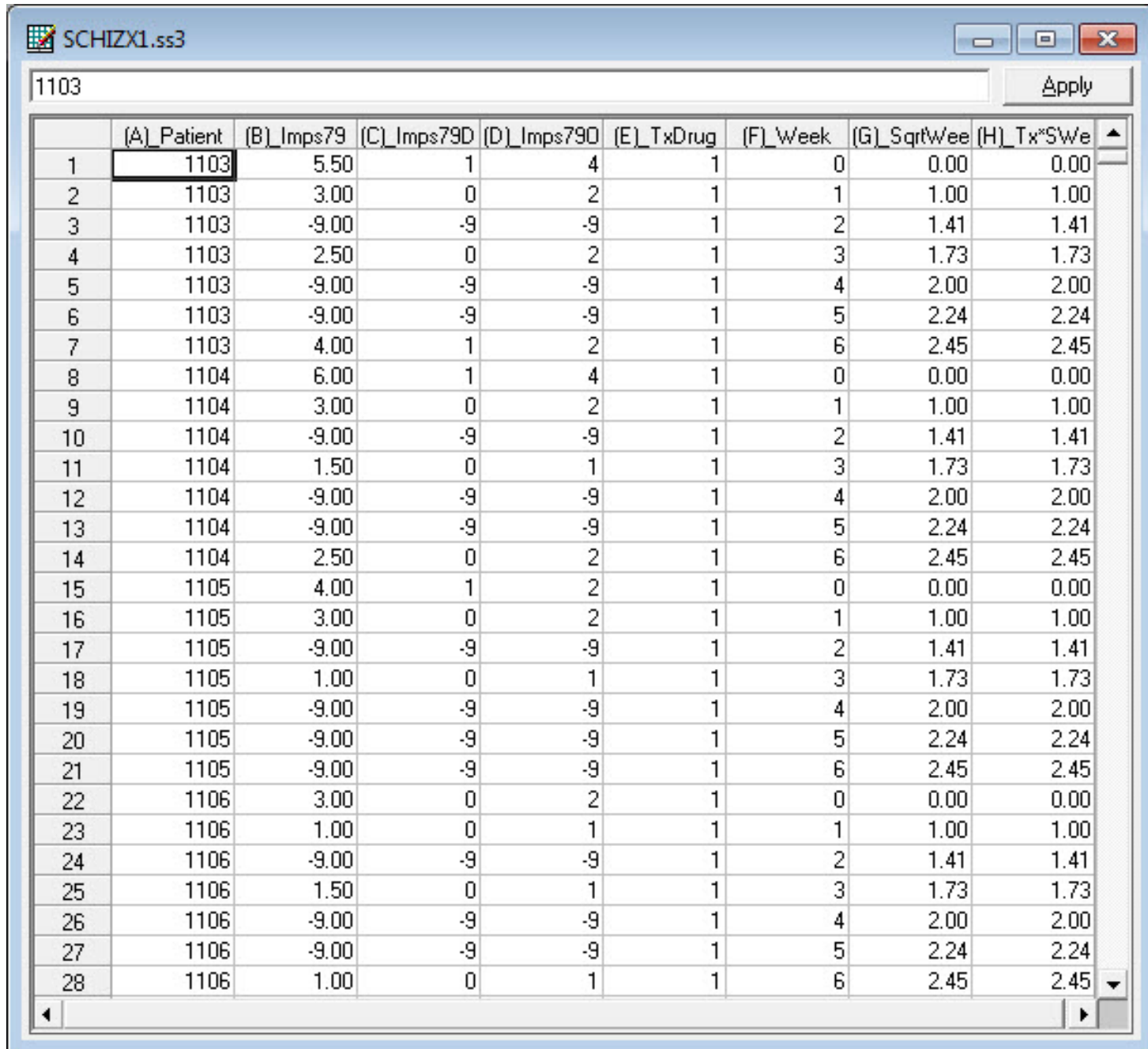
- Under “File” click on “Open Spreadsheet”



- Open C:\SuperMixEn Examples\Workshop\Binary\SCHIZX1.ss3
(or C:\SuperMixEn Student Examples\Workshop\Binary\SCHIZX1.ss3)



C:\SuperMixEn Examples\Workshop\Binary\SCHIZX1.ss3



SCHIZX1.ss3

1103 Apply

	(A)_Patient	(B)_Imps79	(C)_Imps79D	(D)_Imps79D	(E)_TxDrug	(F)_Week	(G)_SqrtWee	(H)_Tx*SWe
1	1103	5.50	1	4	1	0	0.00	0.00
2	1103	3.00	0	2	1	1	1.00	1.00
3	1103	-9.00	-9	-9	1	2	1.41	1.41
4	1103	2.50	0	2	1	3	1.73	1.73
5	1103	-9.00	-9	-9	1	4	2.00	2.00
6	1103	-9.00	-9	-9	1	5	2.24	2.24
7	1103	4.00	1	2	1	6	2.45	2.45
8	1104	6.00	1	4	1	0	0.00	0.00
9	1104	3.00	0	2	1	1	1.00	1.00
10	1104	-9.00	-9	-9	1	2	1.41	1.41
11	1104	1.50	0	1	1	3	1.73	1.73
12	1104	-9.00	-9	-9	1	4	2.00	2.00
13	1104	-9.00	-9	-9	1	5	2.24	2.24
14	1104	2.50	0	2	1	6	2.45	2.45
15	1105	4.00	1	2	1	0	0.00	0.00
16	1105	3.00	0	2	1	1	1.00	1.00
17	1105	-9.00	-9	-9	1	2	1.41	1.41
18	1105	1.00	0	1	1	3	1.73	1.73
19	1105	-9.00	-9	-9	1	4	2.00	2.00
20	1105	-9.00	-9	-9	1	5	2.24	2.24
21	1105	-9.00	-9	-9	1	6	2.45	2.45
22	1106	3.00	0	2	1	0	0.00	0.00
23	1106	1.00	0	1	1	1	1.00	1.00
24	1106	-9.00	-9	-9	1	2	1.41	1.41
25	1106	1.50	0	1	1	3	1.73	1.73
26	1106	-9.00	-9	-9	1	4	2.00	2.00
27	1106	-9.00	-9	-9	1	5	2.24	2.24
28	1106	1.00	0	1	1	6	2.45	2.45

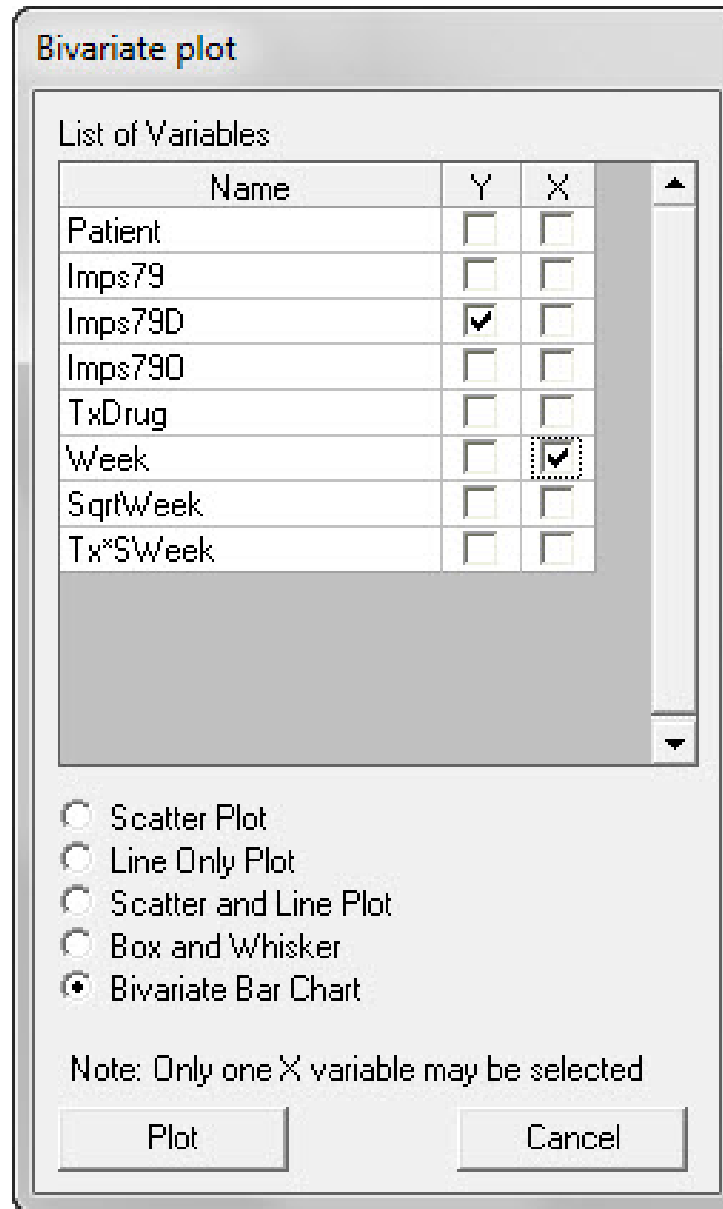
Select Imps79D column, then “Edit” > “Set Missing Value”

The screenshot shows a data editor window titled "SCHIZX1.ss3" with a table of 28 rows and 8 columns. The columns are labeled (A)_Patient, (B)_Imps79, (C)_Imps79D, (D)_Imps79D, (E)_TxDrug, (F)_Week, (G)_SqrtWee, and (H)_Tx*SWe. The data in the table is as follows:

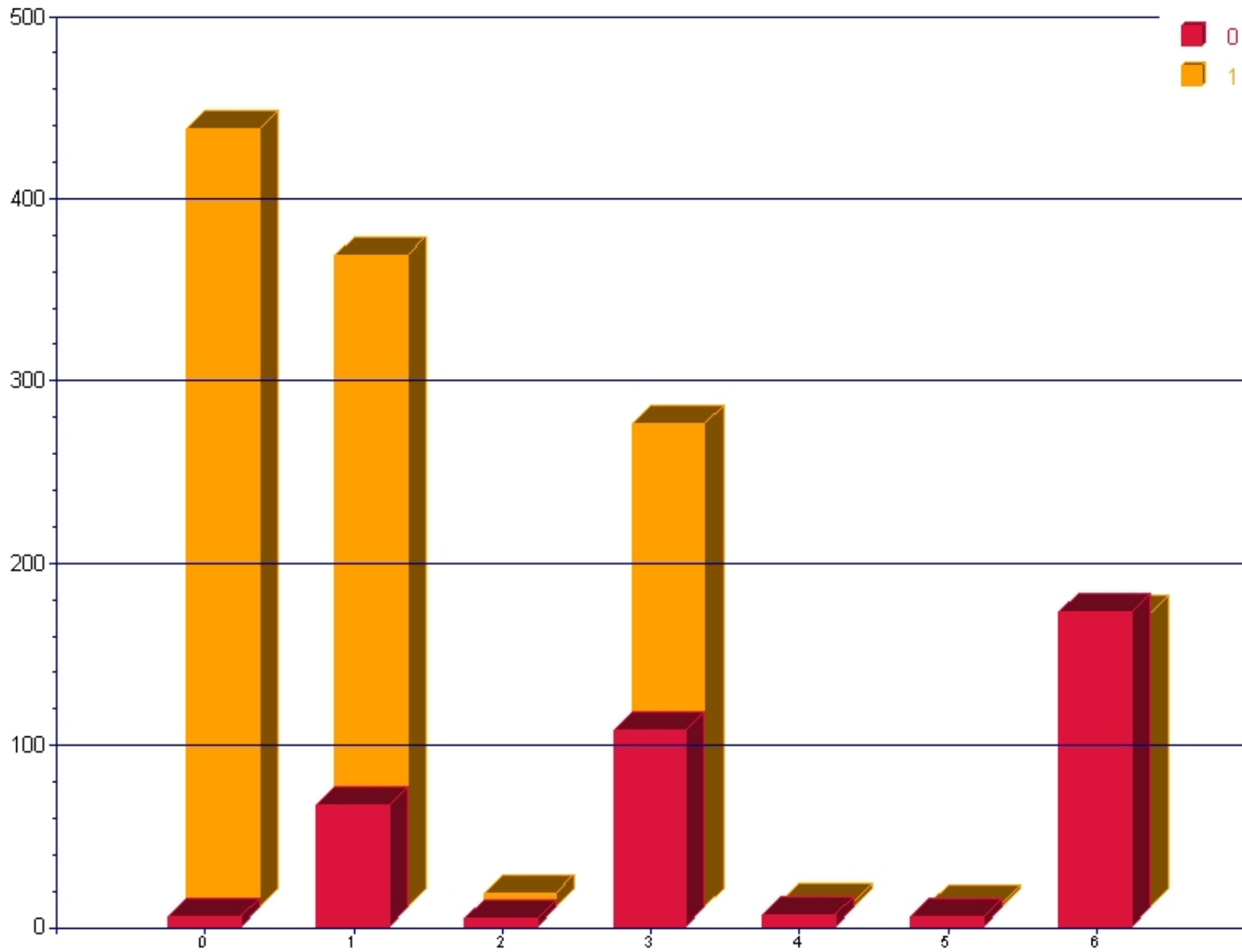
	(A)_Patient	(B)_Imps79	(C)_Imps79D	(D)_Imps79D	(E)_TxDrug	(F)_Week	(G)_SqrtWee	(H)_Tx*SWe
1	1103	5.50	1	4	1	0	0.00	0.00
2	1103	3.00	0	2	1	1	1.00	1.00
3	1103	-9.00	-9	-9	1	2	1.41	1.41
4	1103	2.50	0	2	1	3	1.73	1.73
5	1103	-9.00	-9	-9	1	4	2.00	2.00
6	1103	-9.00	-9	-9	1	5	2.24	2.24
7	1103	4.00	1	2	1	6	2.45	2.45
8	1104	6.00	1	4	1	0	0.00	0.00
9	1104	3.00	0	2	1	1	1.00	1.00
10	1104	-9.00	-9	-9	1	2	1.41	1.41
11	1104	1.50	0	1	1	3	1.73	1.73
12	1104	-9.00	-9	-9	1	4	2.00	2.00
13	1104	-9.00	-9	-9	1	5	2.24	2.24
14	1104	2.50	0	2	1	6	2.45	2.45
15	1105	4.00	1	2	1	0	0.00	0.00
16	1105	3.00	0	2	1	1	1.00	1.00
17	1105	-9.00	-9	-9	1	2	1.41	1.41
18	1105	1.00	0	1	1	3	1.73	1.73
19	1105	-9.00	-9	-9	1	4	2.00	2.00
20	1105	-9.00	-9	-9	1	5	2.24	2.24
21	1105	-9.00	-9	-9	1	6	2.45	2.45
22	1106	3.00	0	2	1	0	0.00	0.00
23	1106	1.00	0	1	1	1	1.00	1.00
24	1106	-9.00	-9	-9	1	2	1.41	1.41
25	1106	1.50	0	1	1	3	1.73	1.73
26	1106	-9.00	-9	-9	1	4	2.00	2.00
27	1106	-9.00	-9	-9	1	5	2.24	2.24
28	1106	1.00	0	1	1	6	2.45	2.45

A dialog box is open over the table, showing "Missing Value Code: -9" with "OK" and "Cancel" buttons.

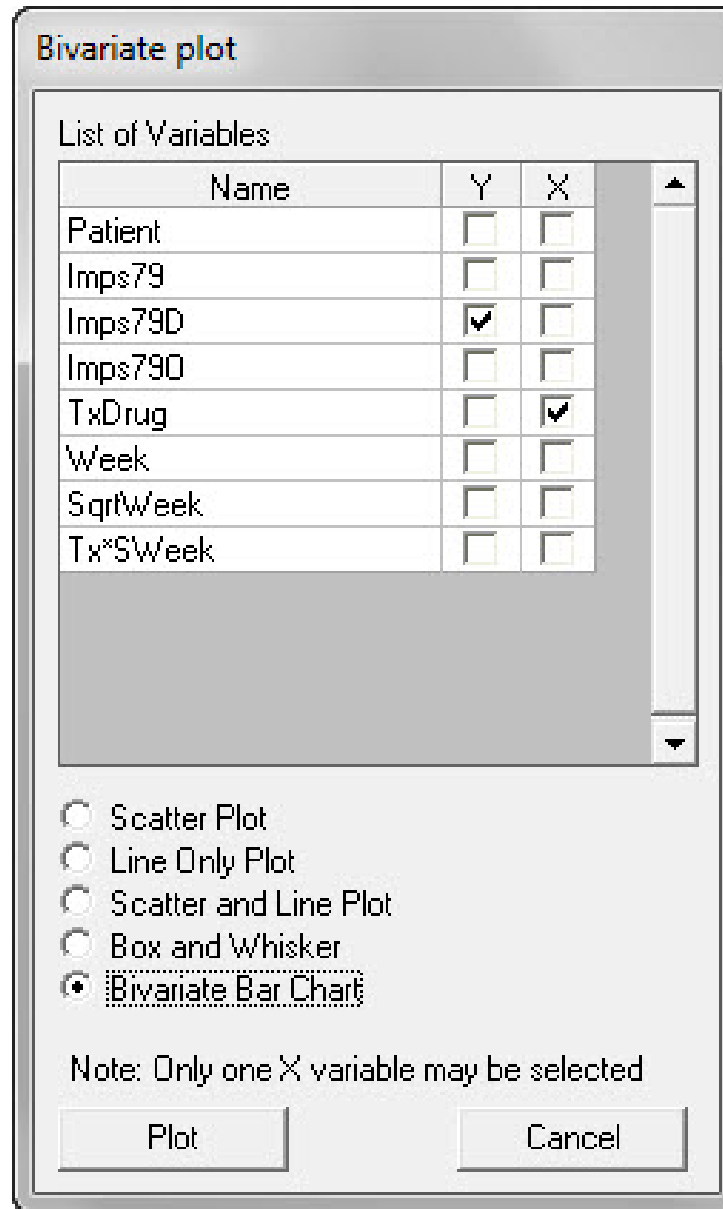
Select “File” > “Data-based Graphs” > “Bivariate”



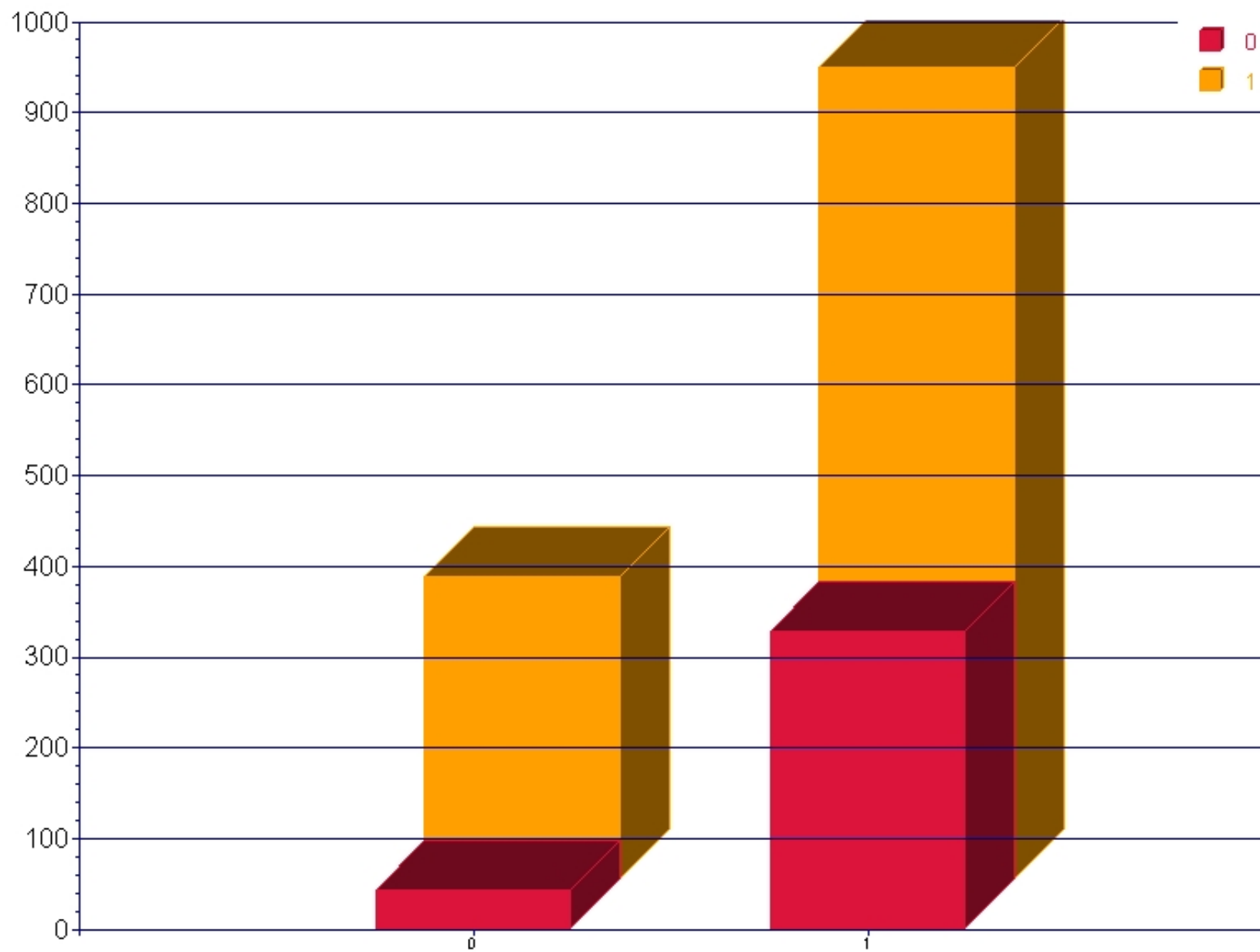
Imps79D vs. Week



Select “File” > “Data-based Graphs” > “Bivariate”



Imps79D vs. TxDrug



Observed proportions \geq “moderately ill”

	<u>week 0</u>	<u>week 1</u>	<u>week 3</u>	<u>week 6</u>
placebo	.98	.91	.89	.71
drug	.99	.82	.66	.42

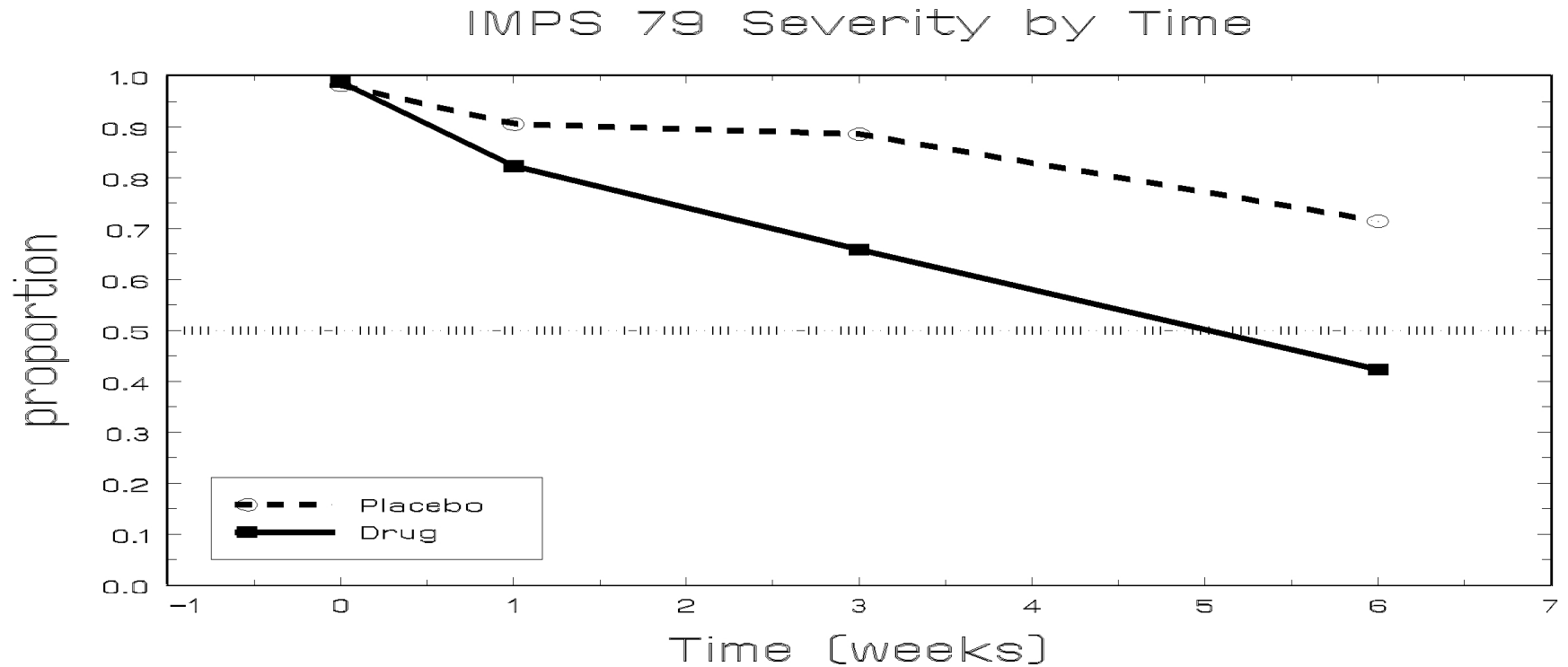
Observed odds \geq “moderately ill”

	<u>week 0</u>	<u>week 1</u>	<u>week 3</u>	<u>week 6</u>
placebo	52.5	9.50	7.70	2.50
drug	80.8	4.63	1.93	.73
<i>ratio</i>	.65	2.05	3.99	3.42

Observed log odds \geq “moderately ill”

	<u>week 0</u>	<u>week 1</u>	<u>week 3</u>	<u>week 6</u>
placebo	3.96	2.25	2.04	.92
drug	4.39	1.53	.66	-.31
<i>difference</i>	-.43	.72	1.38	1.23
exp (odds ratio)	.65	2.05	3.99	3.42

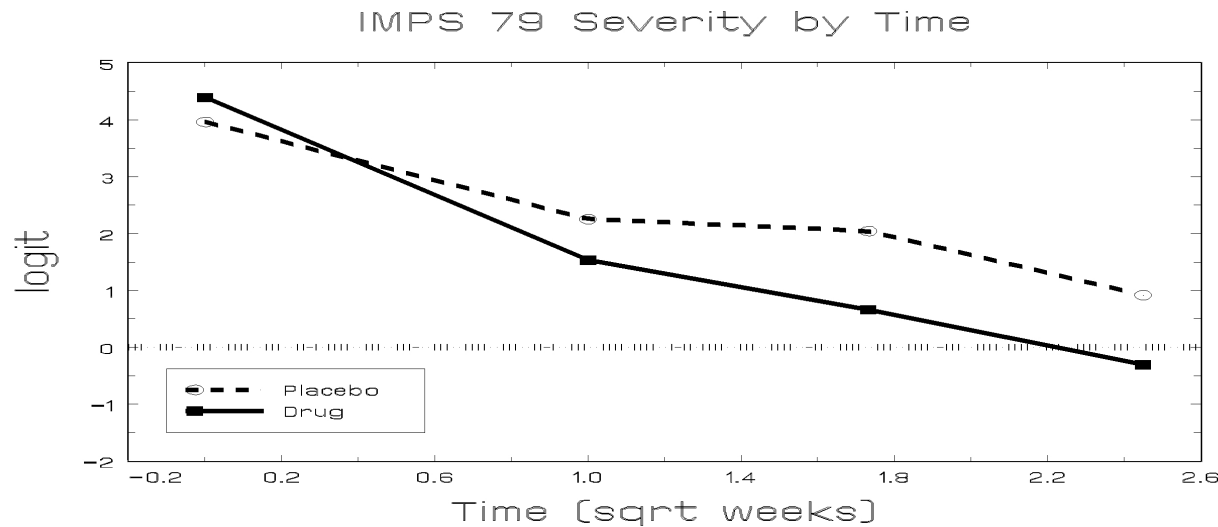
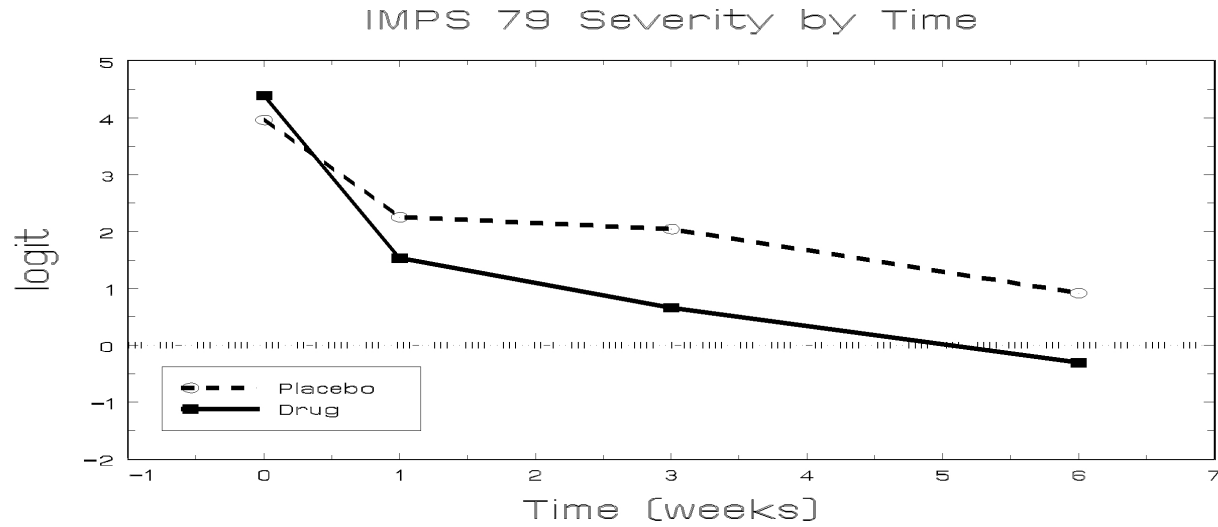
Observed Proportions across Time by Condition



- model is not linear in terms of probabilities

$$P(Y_{ij} = 1 \mid v_{0i}) = \frac{1}{1 + \exp \left[- \left(\mathbf{x}'_{ij} \boldsymbol{\beta} + v_{0i} \right) \right]}$$

Observed Logits across Time by Condition



model is linear in terms of logits: $\log \left[\frac{P(Y_{ij} = 1 | v_{0i})}{1 - P(Y_{ij} = 1 | v_{0i})} \right] = \mathbf{x}'_{ij} \boldsymbol{\beta} + v_{0i}$

Within-Subjects / Between-Subjects components

Within-subjects model - level 1 ($j = 1, \dots, n_i$ obs)

$$\text{logit}_{ij} = b_{0i} + b_{1i}\sqrt{\text{Week}_j}$$

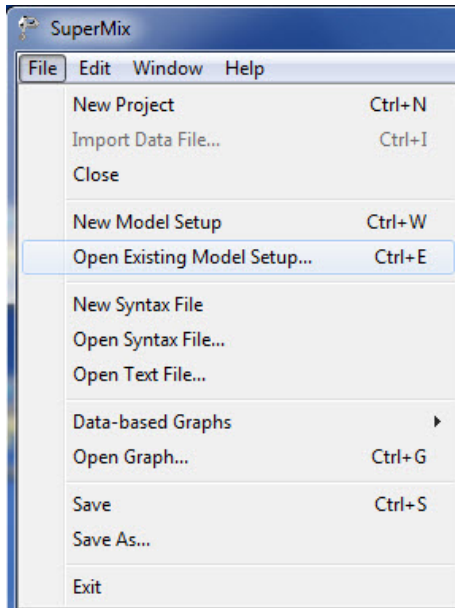
Between-subjects model - level 2 ($i = 1, \dots, N$ subjects)

$$b_{0i} = \beta_0 + \beta_2 \text{Grp}_i + v_{0i}$$

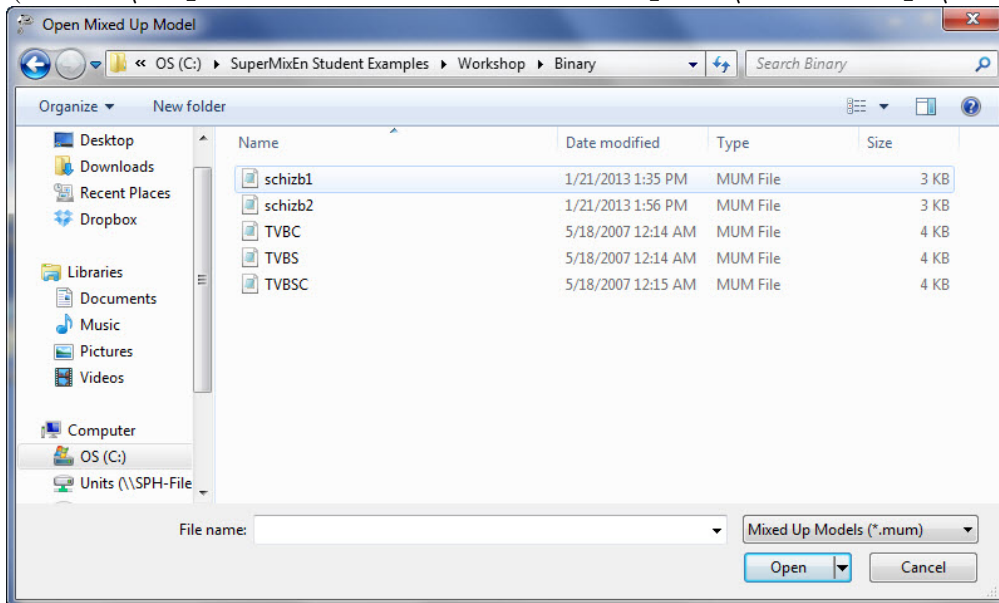
$$b_{1i} = \beta_1 + \beta_3 \text{Grp}_i$$

$$v_{0i} \sim \mathcal{NID}(0, \sigma_v^2)$$

Under “File” click on “Open Existing Model Setup”



Open C:\SuperMixEn Examples\Workshop\Binary\schizb1.mum
(or C:\SuperMixEn Student Examples\Workshop\Binary\schizb1.mum)



Note “Dependent Variable Type” should be “binary”

Model Setup: schizb1.mum

Configuration | Variables | Starting Values | Patterns | Advanced | Linear Transforms

Title 1: Schiz BINARY outcome

Title 2: random intercept model

Dependent Variable Type: **binary**

Level-2 ID: Patient

Dependent Variable: Imps79D

Level-3 ID:

Write Bayes Estimates: no

Convergence Criterion: 0.0001

Number of Iterations: 100

Categories:

	Value
1	0
2	1

Missing Values Present: true

Perform Crosstabulation: no

Missing Value for the Dependent Var: -9.0

Global Missing Value: -9.0

Output Type: standard

Select the form of the dependent variable. The options on the screens will change as required.

Model Setup: schizb1.mum

Configuration | **Variables** | Starting Values | Patterns | Advanced | Linear Transforms

Available	E	2
Patient	<input type="checkbox"/>	<input type="checkbox"/>
Imps79	<input type="checkbox"/>	<input type="checkbox"/>
Imps79D	<input type="checkbox"/>	<input type="checkbox"/>
Imps790	<input type="checkbox"/>	<input type="checkbox"/>
TxDrug	<input checked="" type="checkbox"/>	<input type="checkbox"/>
Week	<input type="checkbox"/>	<input type="checkbox"/>
SqrtWeek	<input checked="" type="checkbox"/>	<input type="checkbox"/>
Tx*SWeek	<input checked="" type="checkbox"/>	<input type="checkbox"/>

Explanatory Variables

TxDrug
SqrtWeek
Tx*SWeek

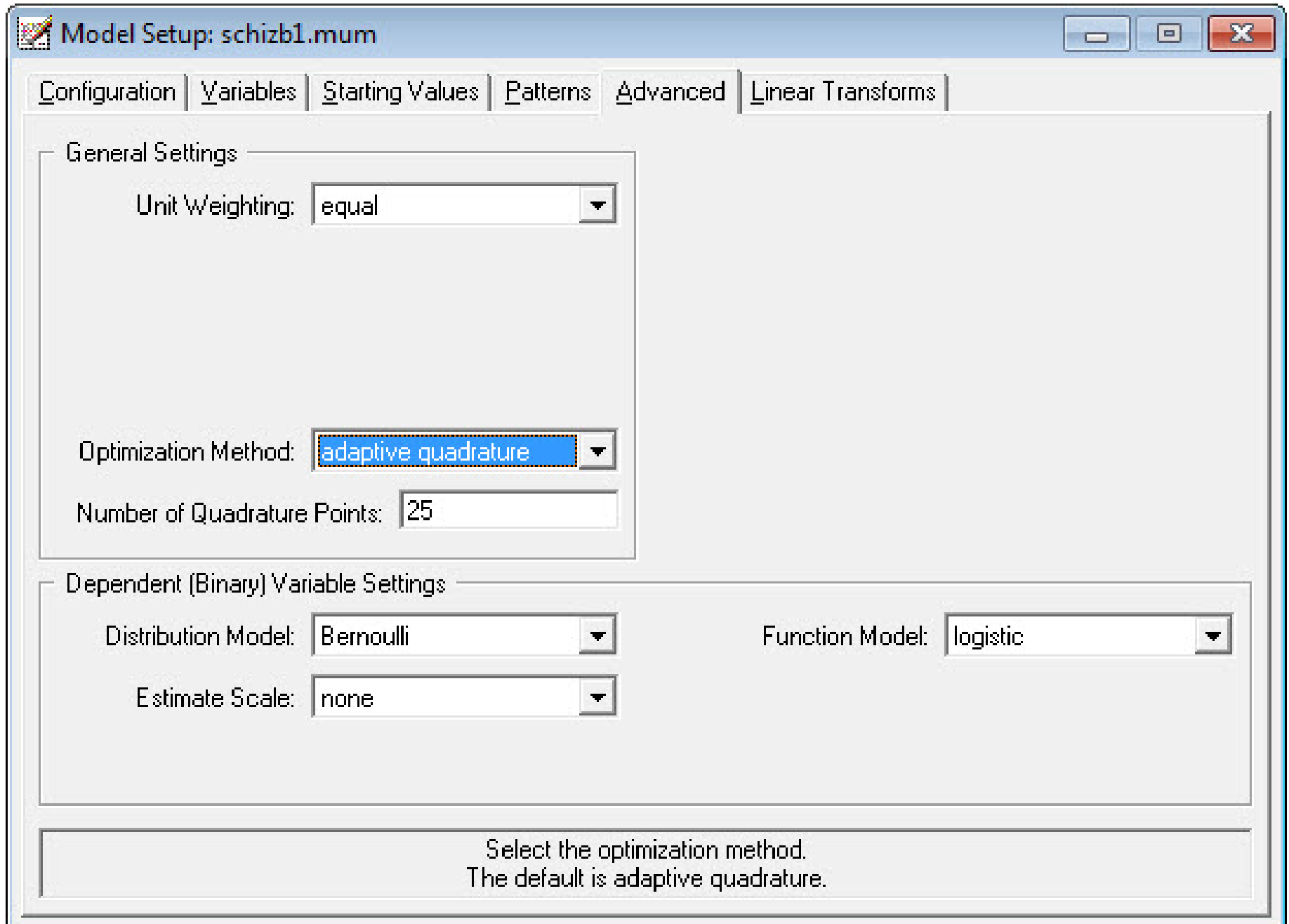
L-2 Random Effects

Include Intercept

Include Intercept

Use the arrow keys or click on the desired tab to select the category of interest for the model.

Note “Optimization Method” should be “adaptive quadrature”



schizb1.out

```

| Schiz BINARY outcome |
| random intercept model |
=====

```

Model and Data Descriptions

```

Sampling Distribution           = Bernoulli
Link Function                  = Logistic
PROB(Success)= 1.0/[1.0+EXP(-ETA)]

```

Number of Level-2 Units 437
Number of Level-1 Units 1603
Number of Level-1 Units per Level-2 Unit =

4	4	3	4	4	4	4	4	4	3	4	4
4	2	3	4	3	4	3	4	4	4	3	3
2	4	4	4	4	4	3	4	4	4	4	4
4	4	4	4	2	3	4	3	4	4	4	3
4	4	2	2	4	5	4	2	4	4	3	4
4	3	2	3	4	4	4	4	4	4	2	4
4	4	5	4	4	2	2	4	2	4	4	3
3	4	4	4	4	4	4	4	4	3	3	4
2	3	4	4	4	2	5	3	4	4	2	4
4	4	2	4	4	4	4	4	4	4	4	4
5	2	4	3	4	4	2	2	4	4	4	4
4	2	4	4	4	4	4	4	4	4	4	4
4	4	4	2	4	4	2	4	4	4	3	4
2	4	4	3	2	3	4	4	3	3	4	3
4	4	4	4	4	4	4	4	4	4	4	4
4	4	2	3	3	5	4	3	4	4	3	2
4	4	4	4	4	3	3	4	4	4	4	4
4	4	4	4	4	4	4	4	4	4	3	4
4	4	4	4	4	2	3	4	4	4	4	4
4	4	4	3	4	4	4	4	4	4	4	4
3	4	4	3	4	4	2	4	4	4	4	2
4	4	4	2	4	4	4	3	3	4	3	4
2	4	4	4	3	3	4	4	4	4	3	3
4	3	4	4	4	4	4	3	4	4	4	4
4	3	3	4	2	4	4	4	4	4	4	4
4	3	4	4	3	3	4	2	4	3	3	3
4	4	4	4	4	4	4	3	2	3	4	4
.

Save As... Close

schizbl.out

Optimization Method: Adaptive Quadrature

Number of quadrature points = 25
 Number of free parameters = 5
 Number of iterations used = 6

-2lnL (deviance statistic) = 1249.73465
 Akaike Information Criterion 1259.73465
 Schwarz Criterion 1286.63281

Estimated regression weights

Parameter	Estimate	Standard Error	z Value	P Value
intercept	5.3851	0.6303	8.5432	0.0000
TxDrug	-0.0247	0.6533	-0.0378	0.9698
SqrtWeek	-1.4996	0.2906	-5.1606	0.0000
Tx*SWeek	-1.0143	0.3338	-3.0385	0.0024

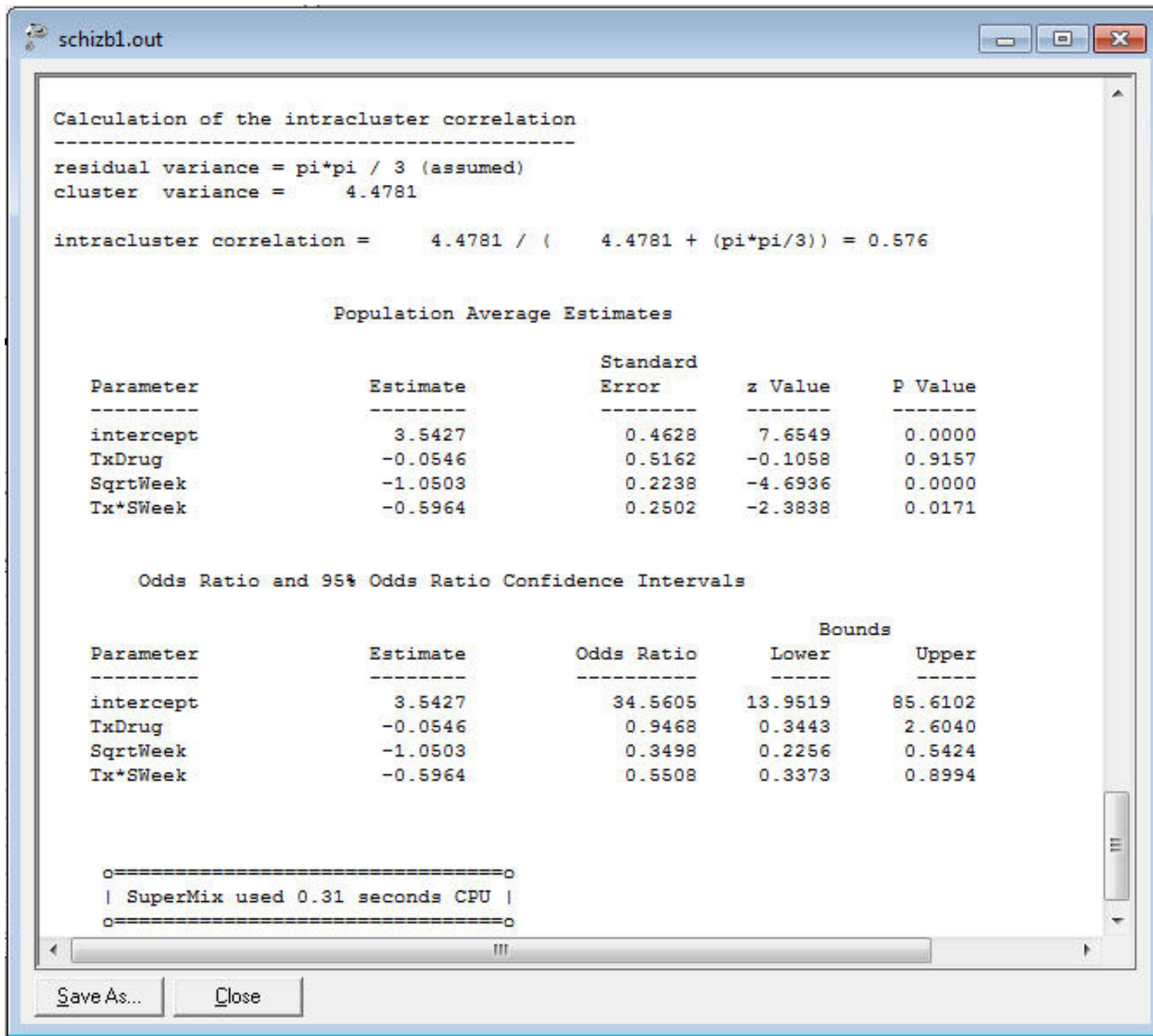
Odds Ratio and 95% Odds Ratio Confidence Intervals

Parameter	Estimate	Odds Ratio	Bounds	
			Lower	Upper
intercept	5.3851	218.1386	63.4125	750.3956
TxDrug	-0.0247	0.9756	0.2711	3.5106
SqrtWeek	-1.4996	0.2232	0.1263	0.3945
Tx*SWeek	-1.0143	0.3627	0.1885	0.6977

Estimated level 2 variances and covariances

Parameter	Estimate	Standard Error	z Value	P Value
intercept/intercept	4.4781	0.9458	4.7345	0.0000

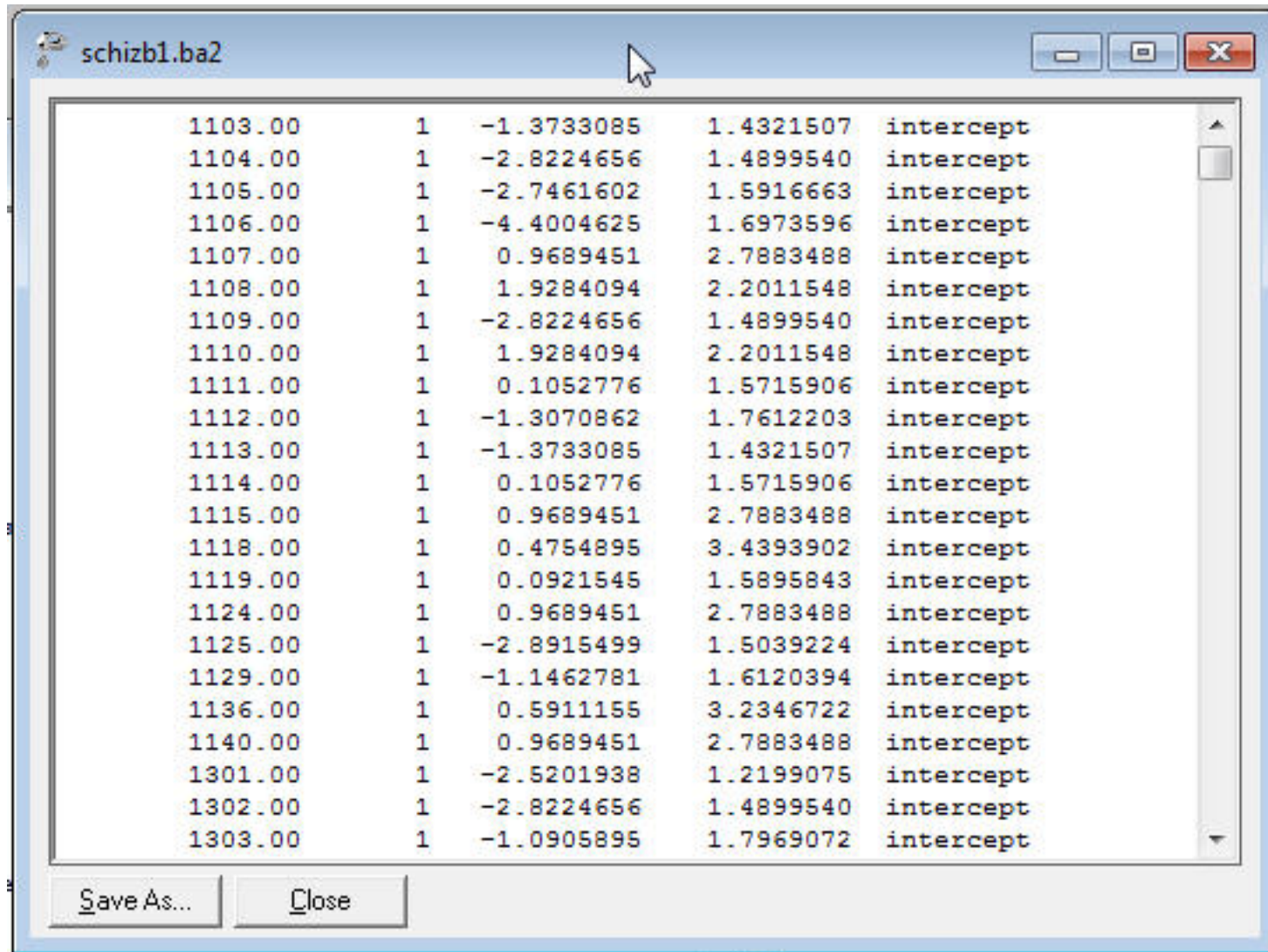
Save As... Close



SuperMix is FAST for full-likelihood estimation of non-normal models, and up to three level models

Empirical Bayes Estimates of Random Effects

Select “Analysis” > “View Level-2 Bayes Results”



ID	random effect number	estimate	variance	name
1103.00	1	-1.3733085	1.4321507	intercept
1104.00	1	-2.8224656	1.4899540	intercept
1105.00	1	-2.7461602	1.5916663	intercept
1106.00	1	-4.4004625	1.6973596	intercept
1107.00	1	0.9689451	2.7883488	intercept
1108.00	1	1.9284094	2.2011548	intercept
1109.00	1	-2.8224656	1.4899540	intercept
1110.00	1	1.9284094	2.2011548	intercept
1111.00	1	0.1052776	1.5715906	intercept
1112.00	1	-1.3070862	1.7612203	intercept
1113.00	1	-1.3733085	1.4321507	intercept
1114.00	1	0.1052776	1.5715906	intercept
1115.00	1	0.9689451	2.7883488	intercept
1118.00	1	0.4754895	3.4393902	intercept
1119.00	1	0.0921545	1.5895843	intercept
1124.00	1	0.9689451	2.7883488	intercept
1125.00	1	-2.8915499	1.5039224	intercept
1129.00	1	-1.1462781	1.6120394	intercept
1136.00	1	0.5911155	3.2346722	intercept
1140.00	1	0.9689451	2.7883488	intercept
1301.00	1	-2.5201938	1.2199075	intercept
1302.00	1	-2.8224656	1.4899540	intercept
1303.00	1	-1.0905895	1.7969072	intercept

ID, random effect number, estimate, variance, name

Close output, select “File” > “Model-based Graphs” > “Equations”

Plot Equations for Outcome Variable

List of Variables

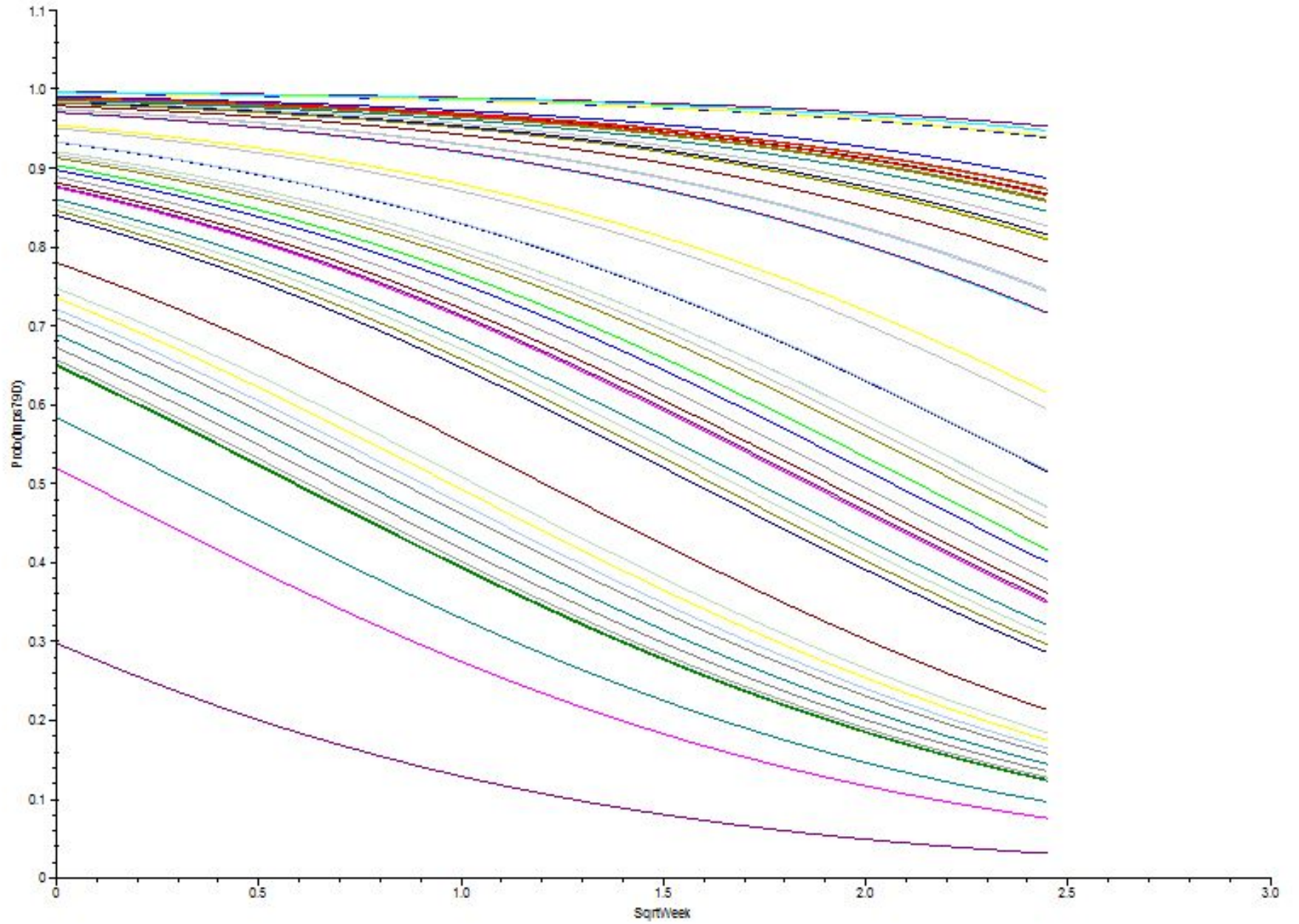
Name	Predictor	Group	Mark
intercept	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
TxDrug	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
SqrtWeek	<input checked="" type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
Tx*S/week	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
Patient		<input type="checkbox"/>	<input checked="" type="checkbox"/>

Remaining predictors fixed at 0
 Remaining predictors fixed at their means
 Plot linear regression model

Note: Only one variable may be selected for grouping and only one for marking.

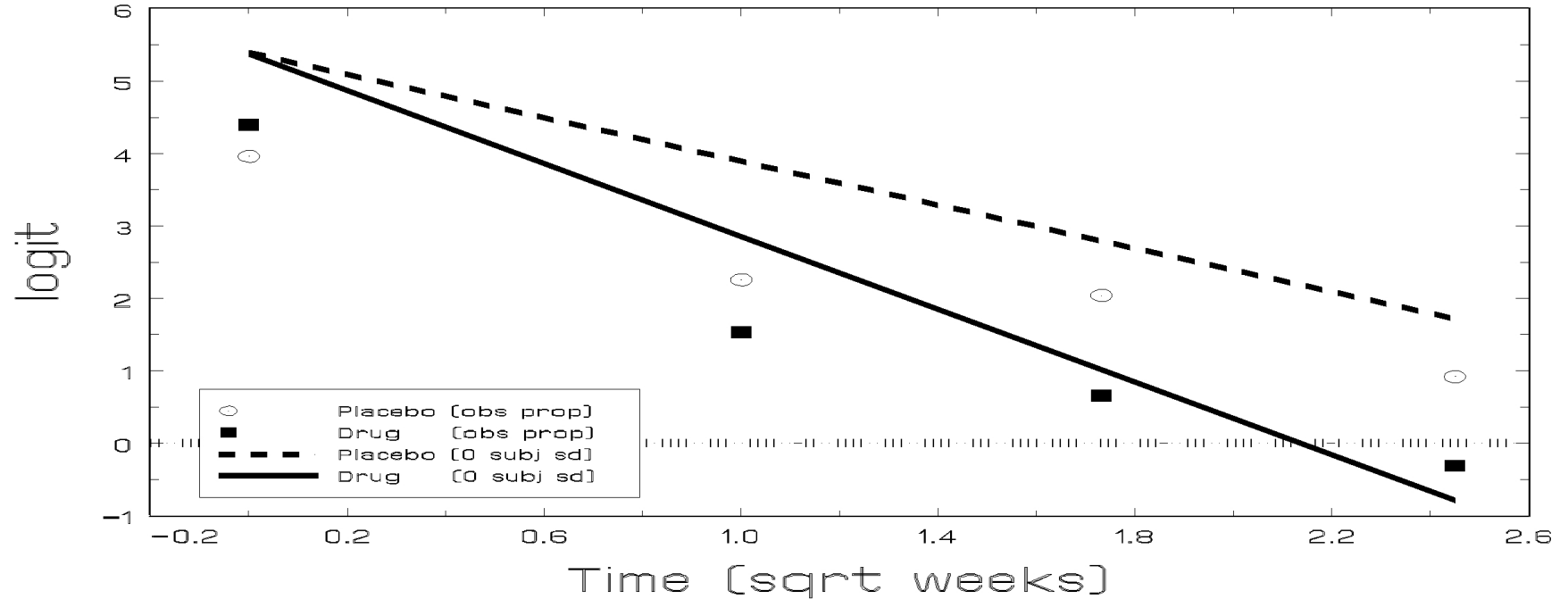
Plot Cancel

Prob(Imps79D) vs. SqrtWeek



Estimated (subject-specific) Logits across Time by Condition: *random-intercepts model*

Random Intercepts Logistic Model



$$\log \left[\frac{P(Y_{ij} = 1 \mid v_{0i})}{1 - P(Y_{ij} = 1 \mid v_{0i})} \right] = 5.39 - .025 D_i - 1.50 T_j - 1.01 (D_i \times T_j) + v_{0i}$$

$$v_{0i} \sim \mathcal{NID}(0, \hat{\sigma}_v^2 = 4.48)$$

$\hat{\beta}$ change in (conditional) logit due to \mathbf{x} for subjects with the same value of v_{0i} (the above plot is for $v_{0i} = 0$)

Random-intercepts Logistic Regression

$$\text{logit}_{ij} = \mathbf{x}'_{ij}\boldsymbol{\beta} + v_{0i}$$

- every subject has their own propensity for response (v_{0i})
- the influence of covariates \mathbf{x} is determined controlling (or adjusting) for the subject effect
- the covariance structure, or dependency, of the repeated observations is explicitly modeled

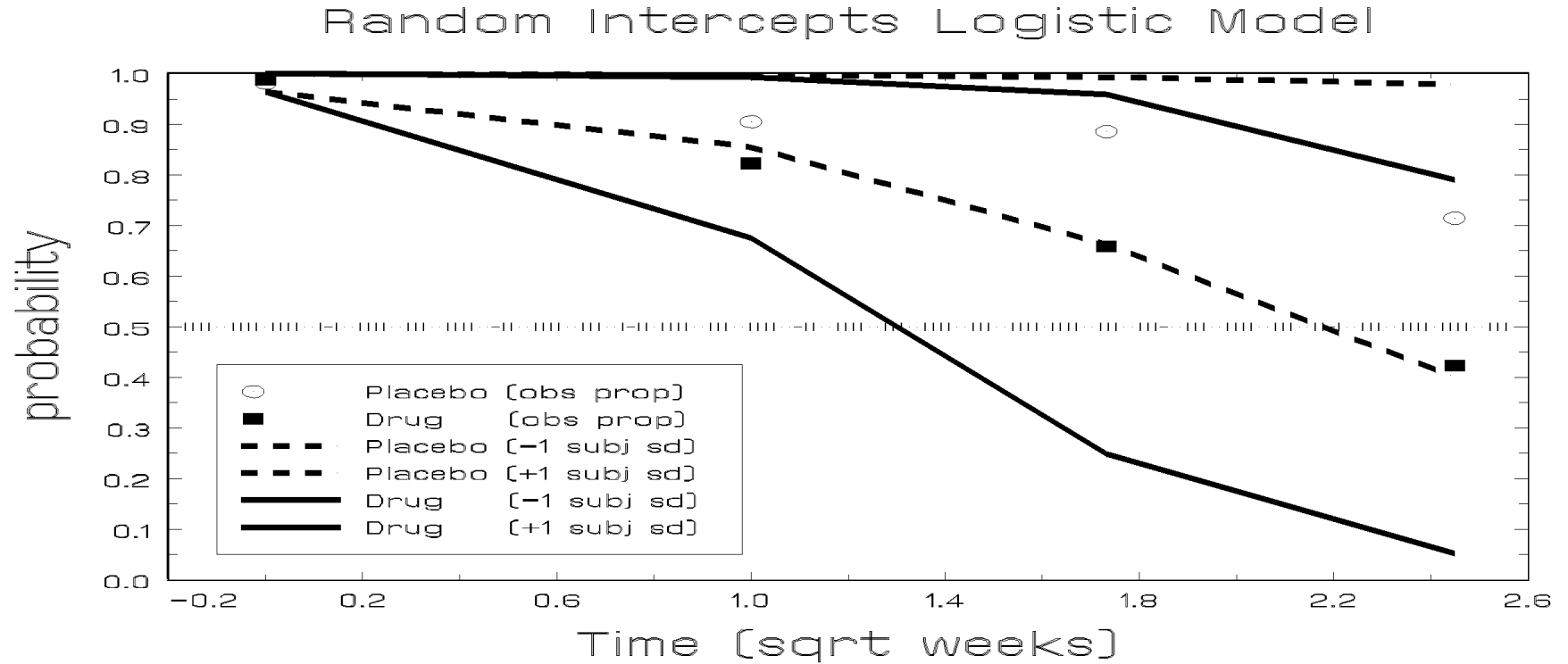
β_0 = log odds of response for a typical subject with $\mathbf{x} = 0$ and $\nu_{0i} = 0$

β = log odds ratio for response associated with unit changes in \mathbf{x} for the same subject value ν_{0i}
* referred to as “subject-specific”
* how a *subject's* response probability depends on \mathbf{x}

σ_ν^2 = degree of heterogeneity across subjects in the probability of response not attributable to \mathbf{x}

- most useful when the objective is to make inference about *subjects* rather than the population average
- interest is in the heterogeneity of subjects

Estimated Subject-Specific Probabilities



$$P(Y_{ij} = 1 \mid v_{0i}) = \frac{1}{1 + \exp[-(5.39 - .03 D_i - 1.50 T_j - 1.01 D_i T_j + v_{0i})]}$$

$$\text{where } v_{0i} = \begin{cases} -1\sigma_v \\ 1\sigma_v \end{cases} \text{ and } \hat{\sigma}_v = 2.12$$

Instead of the mixed model, consider the following marginal model

$$\log \left[\frac{P(Y_{ij} = 1)}{1 - P(Y_{ij} = 1)} \right] = \mathbf{x}'_{ij} \boldsymbol{\beta}^{pa}$$

- $\boldsymbol{\beta}^{pa}$ have marginal or “population-average” interpretation
- Not conditional on subject random effects
- Estimates from a GEE model are of this type
- $\boldsymbol{\beta}^{ss} \neq \boldsymbol{\beta}^{pa}$ unless random effect variance(s) equal 0 (or $\beta = 0$)

\Rightarrow Can one obtain $\boldsymbol{\beta}^{pa}$ from $\boldsymbol{\beta}^{ss}$?

For a random-intercept model with estimates $\hat{\beta}^{ss}$ and $\hat{\sigma}_v^2$

$$\hat{\beta}^{pa} \approx \hat{\beta}^{ss} / \sqrt{\frac{\hat{\sigma}_v^2 + \pi^2/3}{\pi^2/3}}$$

- $\pi^2/3$ is the variance of the standard logistic distribution
- square-root term on the right-hand side can be viewed as the “marginalization” factor; transforms subject-specific parameters into their population-averaged counterparts
- In a random-intercepts model, the variance is equal across time; marginalization factor is equal across time and is a scalar
- For models with multiple random effects, this is not the case and so there is no simple relationship

- Hedeker, du Toit, Demirtas, Gibbons (2014) describe a general marginalization approach that has been implemented in the update of Supermix to yield Population Average Estimates

```

schizbl.out
-----
Calculation of the intracluster correlation
-----
residual variance = pi*pi / 3 (assumed)
cluster variance = 4.4781

intracluster correlation = 4.4781 / ( 4.4781 + (pi*pi/3)) = 0.576

Population Average Estimates

Parameter      Estimate      Standard      z Value      P Value
-----
intercept      3.5427       0.4628       7.6549      0.0000
TxDrug         -0.0546      0.5162      -0.1058     0.9157
SqrtWeek       -1.0503      0.2238      -4.6936     0.0000
Tx*SWeek       -0.5964      0.2502      -2.3838     0.0171

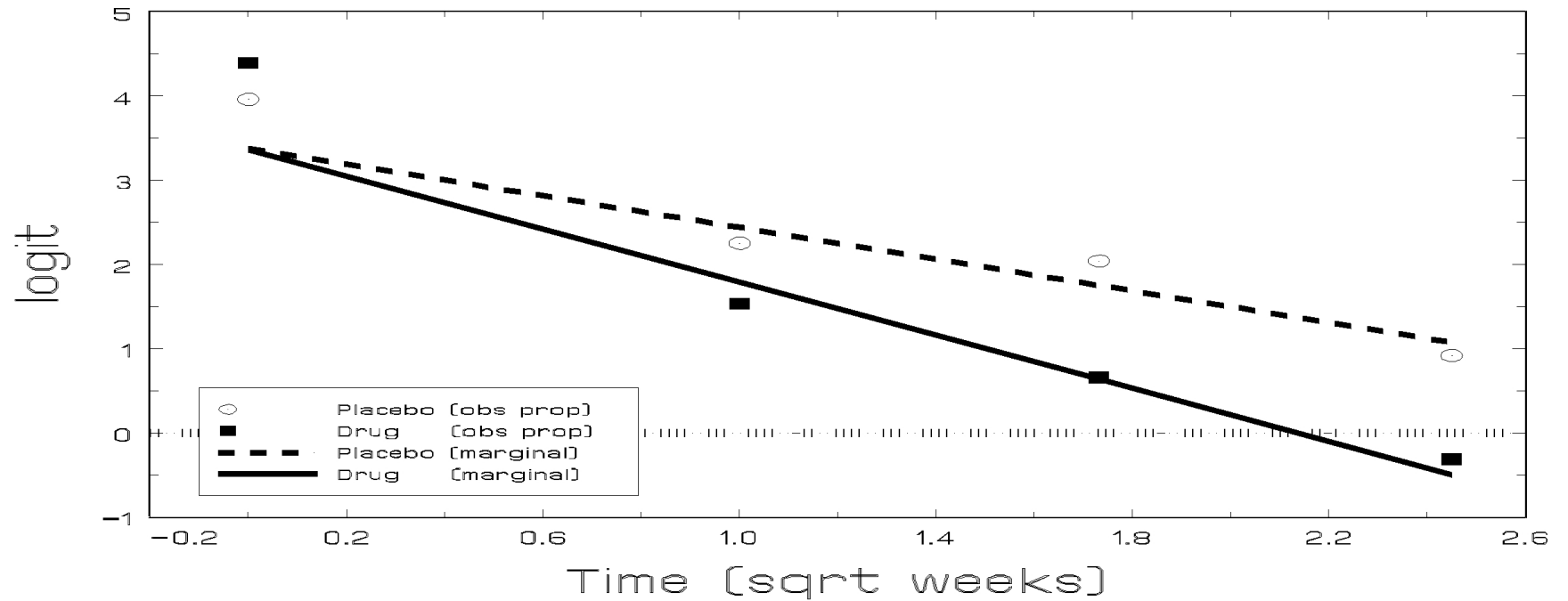
Odds Ratio and 95% Odds Ratio Confidence Intervals

Parameter      Estimate      Odds Ratio      Lower      Upper
-----
intercept      3.5427       34.5605      13.9519     85.6102
TxDrug         -0.0546      0.9468       0.3443     2.6040
SqrtWeek       -1.0503      0.3498       0.2256     0.5424
Tx*SWeek       -0.5964      0.5508       0.3373     0.8994

-----
| SuperMix used 0.31 seconds CPU |
-----
Save As... Close

```

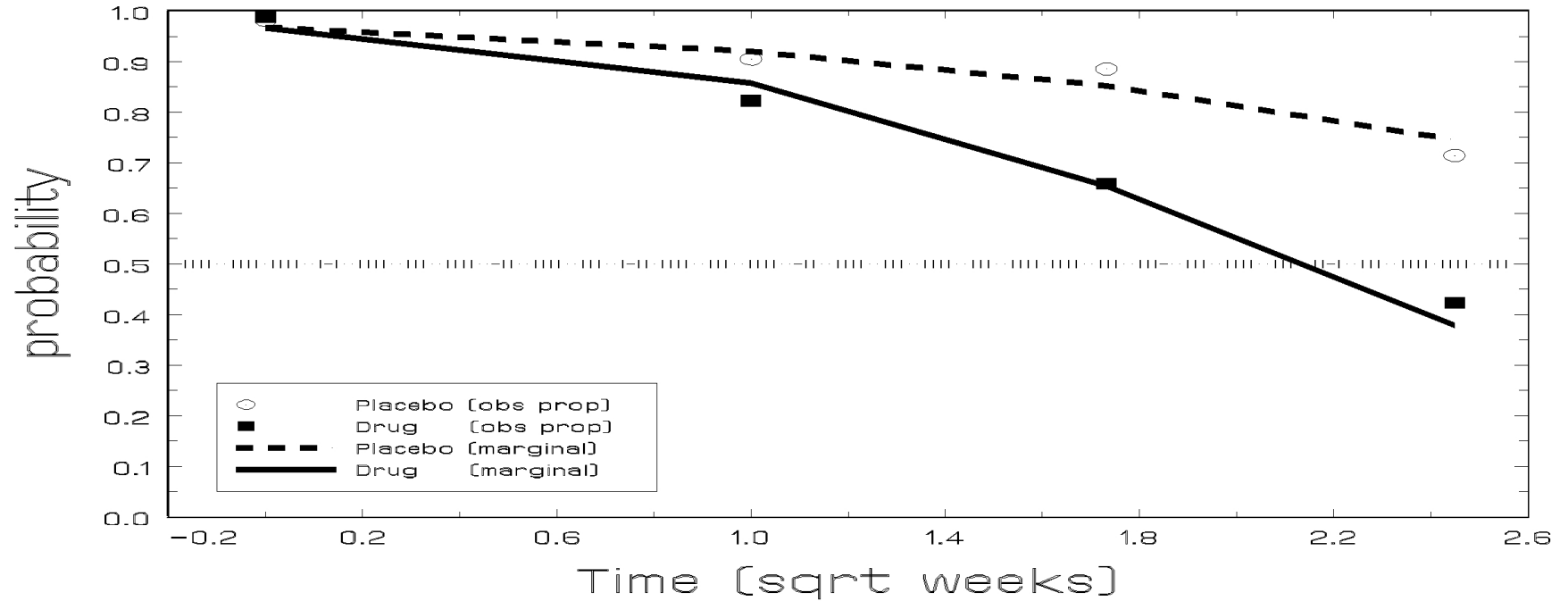
Marginalized Random Intercepts Logistic



$$\log \left[\frac{P(Y_{ij} = 1)}{1 - P(Y_{ij} = 1)} \right] = 3.54 - .055 D_i - 1.05 T_j - .60 (D_i \times T_j)$$

⇒ these are the Population Average Estimates from Supermix

Marginalized Random Intercepts Logistic



$$P(Y_{ij} = 1) = \frac{1}{1 + \exp \left[- \left(3.54 - .055 D_i - 1.05 T_j - .60 D_i T_j \right) \right]}$$

⇒ these are the Population Average Estimates from Supermix

Within-Subjects / Between-Subjects components

Within-subjects model - level 1 ($j = 1, \dots, n_i$ obs)

$$\text{logit}_{ij} = b_{0i} + b_{1i}\sqrt{\text{Week}_j}$$

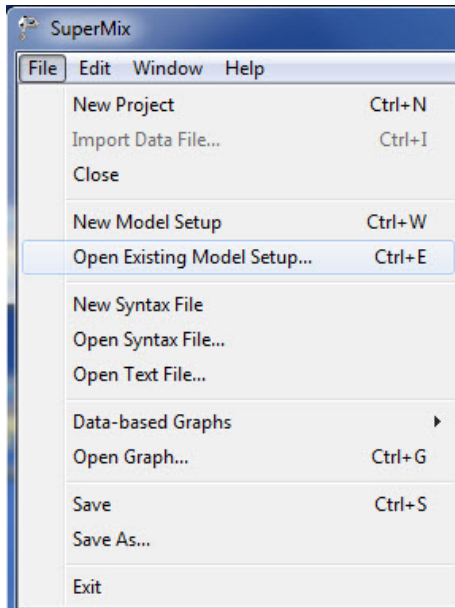
Between-subjects model - level 2 ($i = 1, \dots, N$ subjects)

$$b_{0i} = \beta_0 + \beta_2 \text{Grp}_i + v_{0i}$$

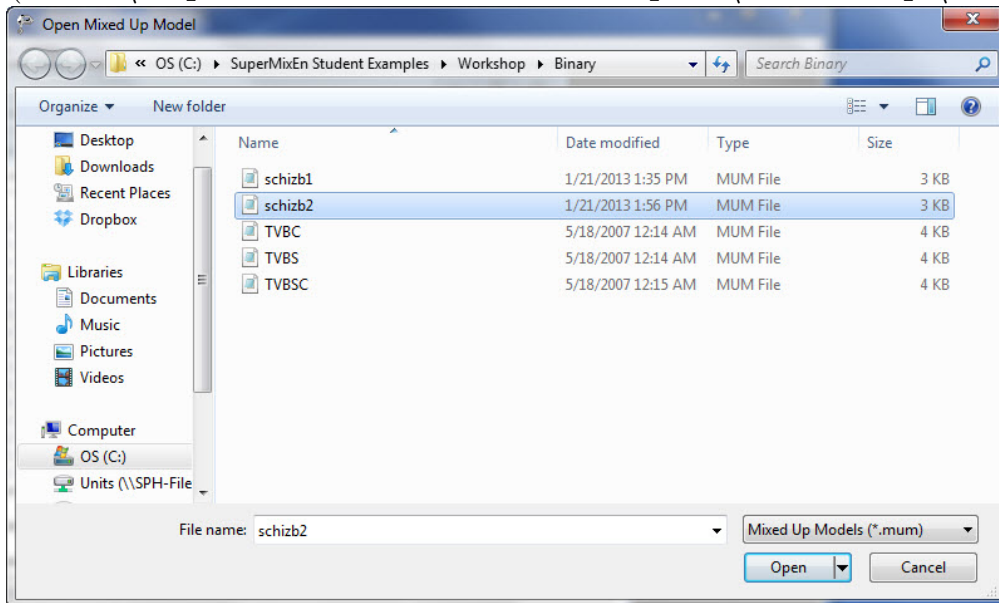
$$b_{1i} = \beta_1 + \beta_3 \text{Grp}_i + v_{1i}$$

$$\mathbf{v}_i \sim \mathcal{NID}(\mathbf{0}, \Sigma_v)$$

Under “File” click on “Open Existing Model Setup”



Open C:\SuperMixEn Examples\Workshop\Binary\schizb2.mum
(or C:\SuperMixEn Student Examples\Workshop\Binary\schizb2.mum)



Note “Dependent Variable Type” should be “binary”

Model Setup: schizb2.mum

Configuration | Variables | Starting Values | Patterns | Advanced | Linear Transforms

Title 1: Schiz BINARY outcome

Title 2: random intercept and trend model

Dependent Variable Type: **binary**

Level-2 IDs: Patient

Dependent Variable: Imps79D

Level-3 IDs:

Categories:

	Value
1	0
2	1

Write Bayes Estimates: no

Convergence Criterion: 0.001

Number of Iterations: 100

Missing Values Present: true

Perform Crosstabulation: no

Missing Value for the Dependent Var: -9.0

Global Missing Value: -9.0

Output Type: standard

Select the form of the dependent variable. The options on the screens will change as required.

SqrtWeek is a level-2 (subject) random effect

Model Setup: schizb2.mum

Configuration | **Variables** | Starting Values | Patterns | Advanced | Linear Transforms

Available	E	2
Patient	<input type="checkbox"/>	<input type="checkbox"/>
Imps79	<input type="checkbox"/>	<input type="checkbox"/>
Imps79D	<input type="checkbox"/>	<input type="checkbox"/>
Imps790	<input type="checkbox"/>	<input type="checkbox"/>
TxDrug	<input checked="" type="checkbox"/>	<input type="checkbox"/>
Week	<input type="checkbox"/>	<input type="checkbox"/>
Sqrt/week	<input checked="" type="checkbox"/>	<input checked="" type="checkbox"/>
Tx*S/week	<input checked="" type="checkbox"/>	<input type="checkbox"/>

Explanatory Variables

- TxDrug
- Tx*S/week
- Sqrt/week

Include Intercept

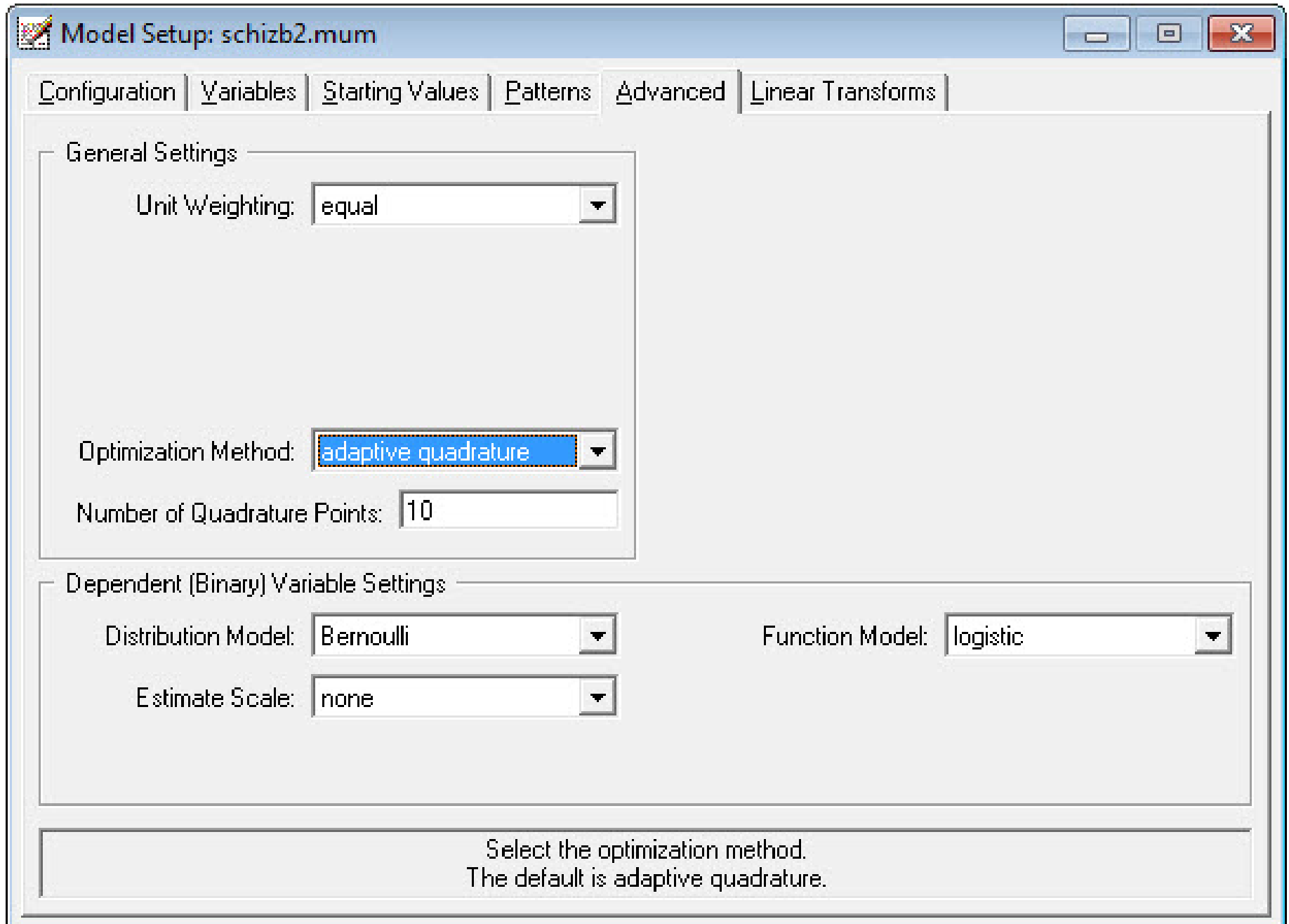
L-2 Random Effects

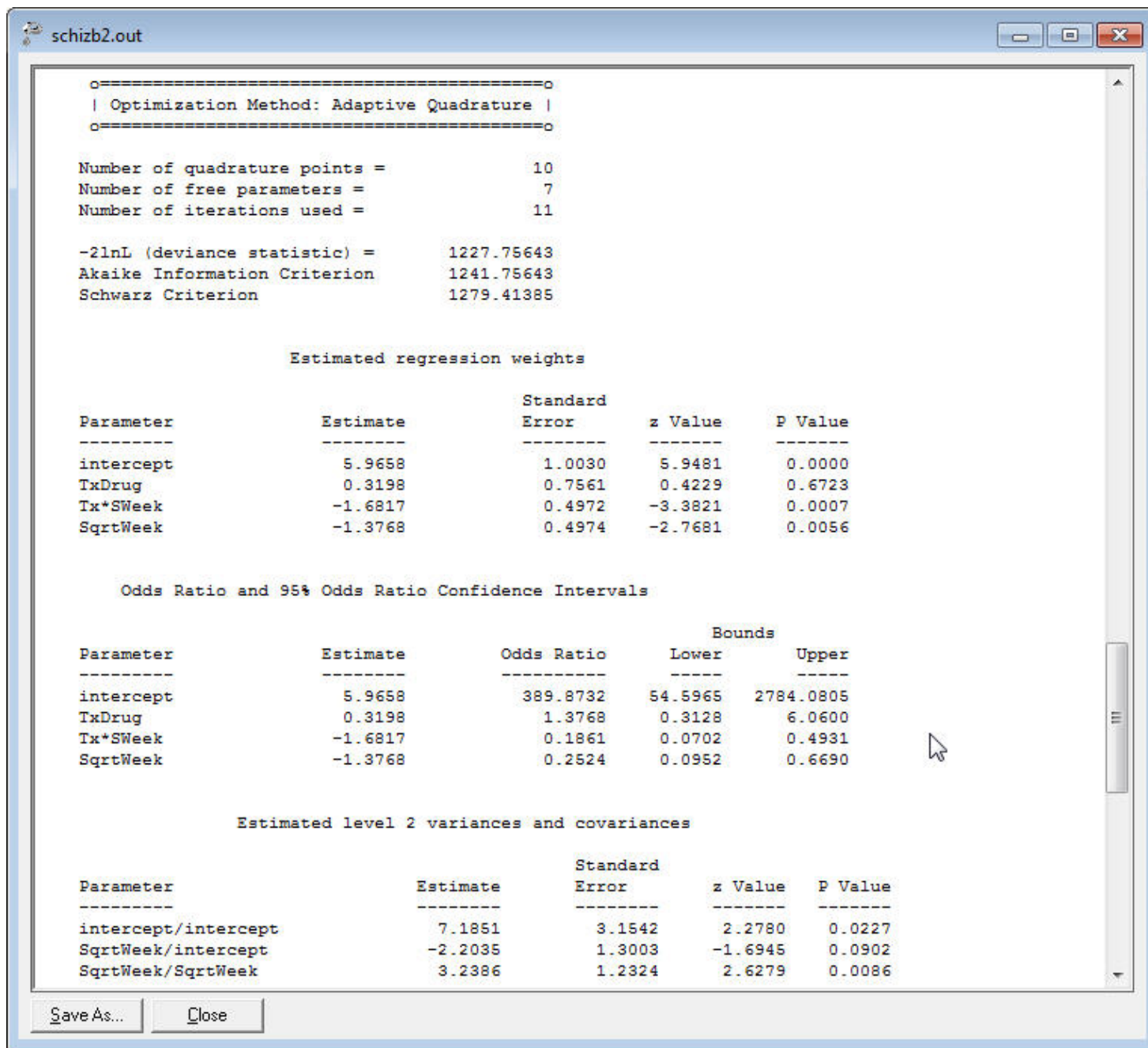
- Sqrt/week

Include Intercept

Use the arrow keys or click on the desired tab to select the category of interest for the model.

Note “Optimization Method” should be “adaptive quadrature”





⇒ Comparing models: $H_0 : \sigma_{v_1}^2 = \sigma_{v_{01}} = 0$; $\chi_2^2 = 1249.73 - 1227.76 = 21.97, p < .001$

```

schizb2.out

intercept      SqrtWeek
intercept      7.185117
SqrtWeek      -2.203484      3.238562

Level 2 correlation matrix

intercept      SqrtWeek
intercept      1.000000
SqrtWeek      -0.456790      1.000000

Population Average Estimates

Parameter      Estimate      Standard      z Value      P Value
-----      -----      -----      -----      -----
intercept      3.4632      0.4745      7.2979      0.0000
TxDrug      0.0472      0.5375      0.0879      0.9300
Tx*SWeek      -0.6609      0.2867      -2.3055      0.0211
SqrtWeek      -0.9955      0.2559      -3.8904      0.0001

Odds Ratio and 95% Odds Ratio Confidence Intervals

Parameter      Estimate      Odds Ratio      Bounds
-----      -----      -----      -----
intercept      3.4632      31.9181      12.5921      80.9052
TxDrug      0.0472      1.0484      0.3656      3.0067
Tx*SWeek      -0.6609      0.5164      0.2944      0.9057
SqrtWeek      -0.9955      0.3696      0.2238      0.6102

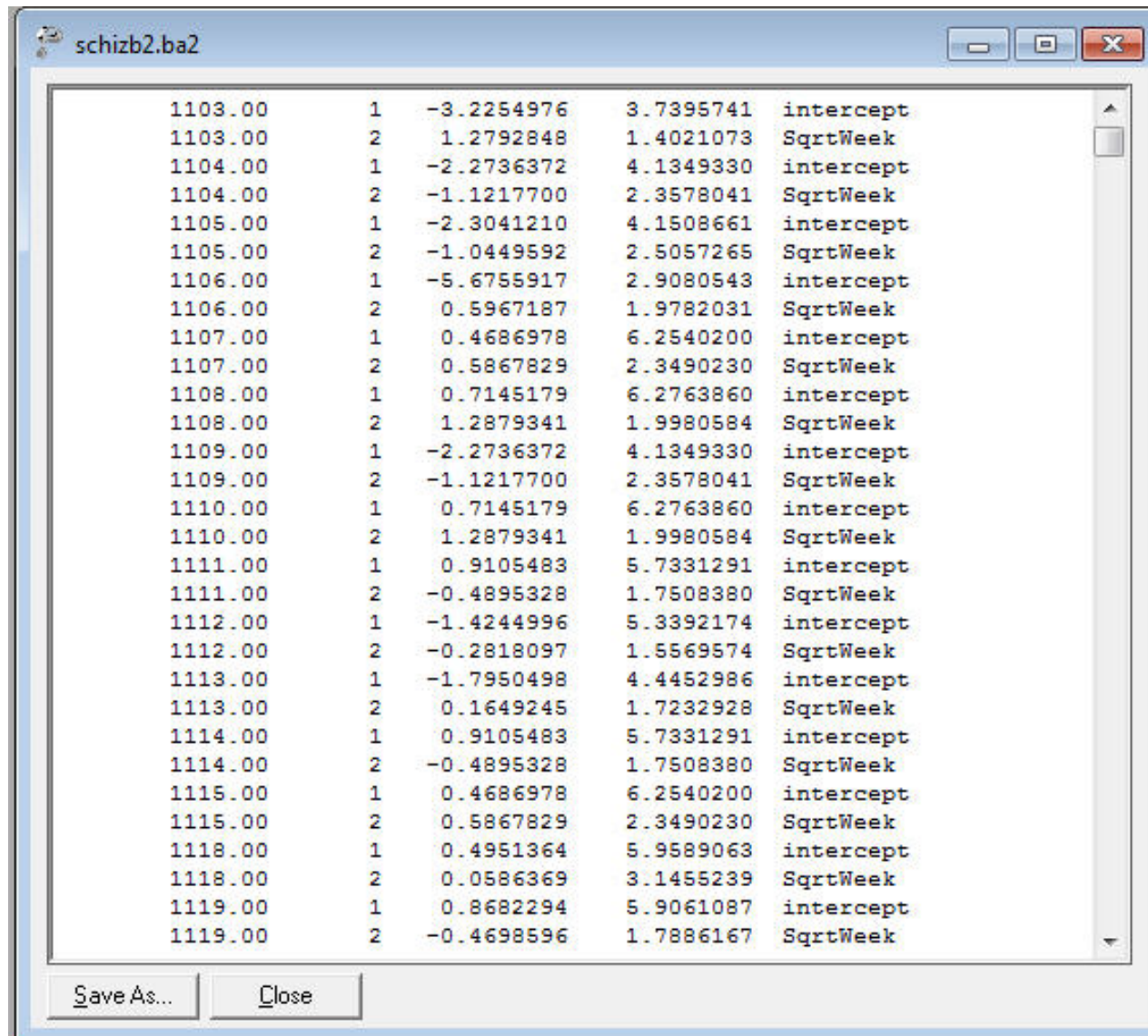
=====
| SuperMix used 1.11 seconds CPU |
=====
Save As... Close

```

⇒ Supermix is FAST for a full-likelihood solution using bivariate numerical integration involving 100 quadrature points!

Empirical Bayes Estimates of Random Effects

Select “Analysis” > “View Level-2 Bayes Results”



ID	random effect number	estimate	variance	name
1103.00	1	-3.2254976	3.7395741	intercept
1103.00	2	1.2792848	1.4021073	SqrtWeek
1104.00	1	-2.2736372	4.1349330	intercept
1104.00	2	-1.1217700	2.3578041	SqrtWeek
1105.00	1	-2.3041210	4.1508661	intercept
1105.00	2	-1.0449592	2.5057265	SqrtWeek
1106.00	1	-5.6755917	2.9080543	intercept
1106.00	2	0.5967187	1.9782031	SqrtWeek
1107.00	1	0.4686978	6.2540200	intercept
1107.00	2	0.5867829	2.3490230	SqrtWeek
1108.00	1	0.7145179	6.2763860	intercept
1108.00	2	1.2879341	1.9980584	SqrtWeek
1109.00	1	-2.2736372	4.1349330	intercept
1109.00	2	-1.1217700	2.3578041	SqrtWeek
1110.00	1	0.7145179	6.2763860	intercept
1110.00	2	1.2879341	1.9980584	SqrtWeek
1111.00	1	0.9105483	5.7331291	intercept
1111.00	2	-0.4895328	1.7508380	SqrtWeek
1112.00	1	-1.4244996	5.3392174	intercept
1112.00	2	-0.2818097	1.5569574	SqrtWeek
1113.00	1	-1.7950498	4.4452986	intercept
1113.00	2	0.1649245	1.7232928	SqrtWeek
1114.00	1	0.9105483	5.7331291	intercept
1114.00	2	-0.4895328	1.7508380	SqrtWeek
1115.00	1	0.4686978	6.2540200	intercept
1115.00	2	0.5867829	2.3490230	SqrtWeek
1118.00	1	0.4951364	5.9589063	intercept
1118.00	2	0.0586369	3.1455239	SqrtWeek
1119.00	1	0.8682294	5.9061087	intercept
1119.00	2	-0.4698596	1.7886167	SqrtWeek

ID, random effect number, estimate, variance, name

Close output, select “File” > “Model-based Graphs” > “Equations”

Plot Equations for Outcome Variable

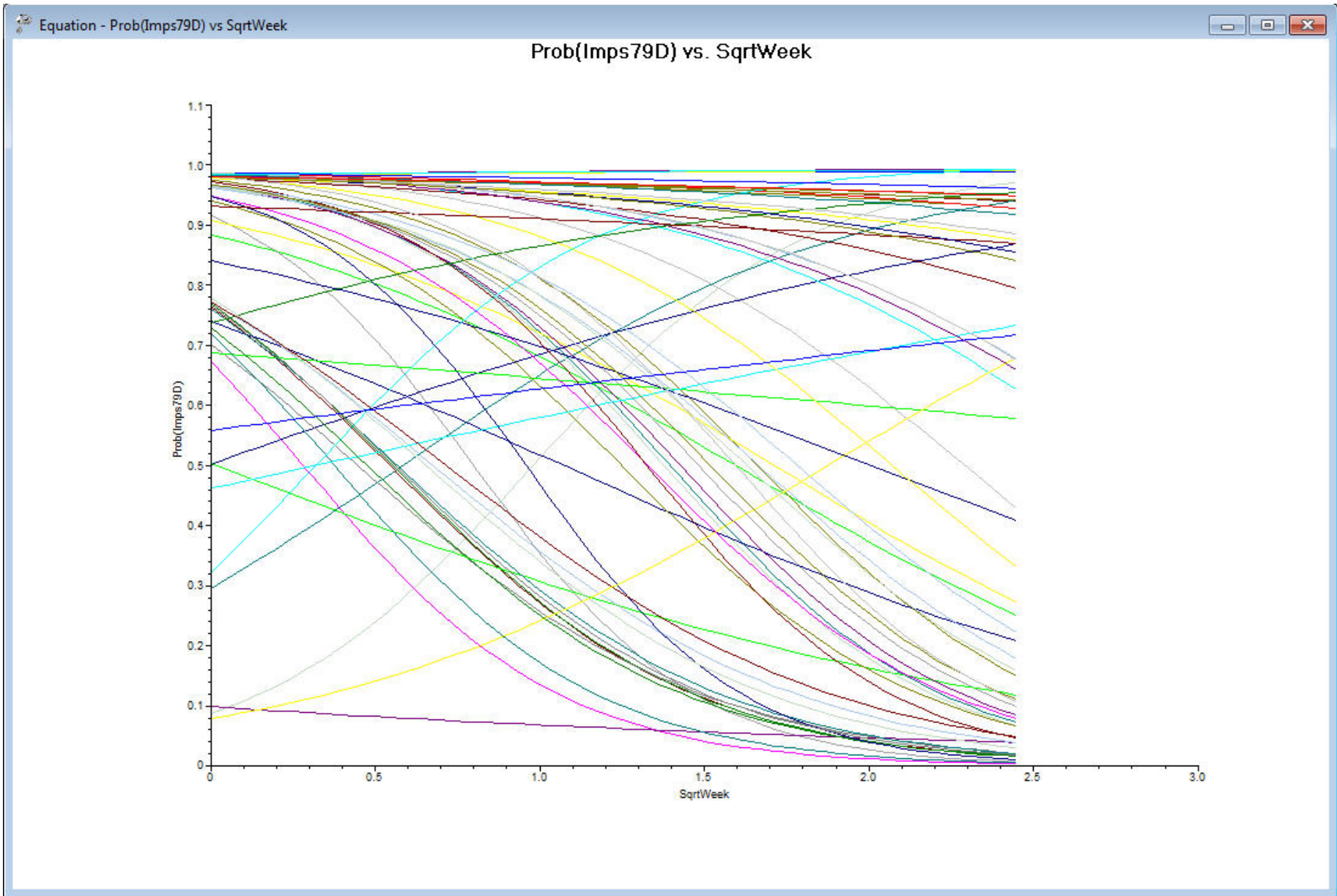
List of Variables

Name	Predictor	Group	Mark
intercept	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
TxDrug	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
SqrtWeek	<input checked="" type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
Tx*S/Week	<input type="checkbox"/>	<input type="checkbox"/>	<input type="checkbox"/>
Patient		<input type="checkbox"/>	<input checked="" type="checkbox"/>

Remaining predictors fixed at 0
 Remaining predictors fixed at their means
 Plot linear regression model

Note: Only one variable may be selected for grouping and only one for marking.

Plot Cancel



Summary - mixed models for binary outcomes

- link functions: logistic, probit, log-log, complementary log-log
- multiple random effects (correlated or independent) for up to 3-level models
- fast full-likelihood estimation using adaptive Gauss-Hermite quadrature
- subject-specific and population-average estimates and inference
- discrete/grouped time survival analysis via person-period dataset
- Advanced > Level-2 (Co)variance Patterns > Unidimensional
 - varying ICC model for MZ/DZ twin pair data: create dummy variables MZ and DZ, specify both as random effects, select “Unidimensional”
 - Item-response theory (IRT) models: create indicator variables for items, specify all as random effects, select “Unidimensional”